

Melting Down: Systemic Financial Instability and the Macroeconomy

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Introduction

- Financial crises are rare disruptive events which challenge macro model building and forecasting
 - Economic history is replete with financial crises, although they still represent rare events in typical macro data samples (for developed economies).
 - Systemic financial crises, i.e. crises in which financial instability spreads widely across the entire financial system, are particularly disruptive, often causing huge losses in output and social welfare.
 - Financial and real shocks generated by a financial crisis seem to give rise to abrupt and unusual (non-linear) changes in economic dynamics, likely caused by certain amplification, propagation and feedback mechanisms activated during times of distress.

Introduction (cont'd)

- Recently growing literature on theoretical macrofinancial models that incorporate financial instability
 - Brunnermeier and Sannikov (AEA 2014), He and Krishnamurthy (RES 2011), Boissay, Collard and Smets (JPE 2015).
- Few empirical contributions on crisis effects
 - Schularick and Taylor (AEA 2012), Giglio, Kelly and Pruitt (NBER 2015), Hubrich and Tetlow (JME 2015).
- This paper presents estimates of a small-scale macrofinancial VAR model for the euro area that
 - allows independent regime shifts in coefficients and shock variances to capture potential non-linearities;
 - integrates features of systemic financial instability by using a composite indicator of systemic stress (CISS) as one of the endogenous model variables.

Preview of main results

- We find strong evidence for non-linearities
 - Both the variances of shocks and the model coefficients characterising the dynamic relationships between the variables undergo recurrent regime changes;
 - Developments in the CISS help interpreting the identified regimes;
 - Effects of shocks in the CISS on economic activity become much larger during high-stress regimes.

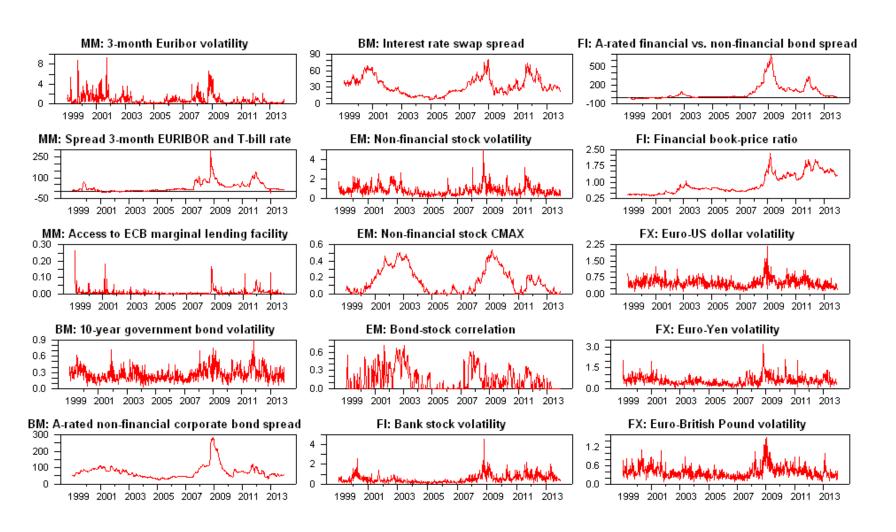
Outline

- 1. Measuring systemic financial instability: The ECB's CISS
- 2. Markov-Switching Bayesian Vector Autoregression (MS-BVAR) model
- 3. Determining and interpreting macroeconomic regimes
- 4. From subprime turmoil to the systemic meltdown: A model tale
- 5. Concluding remarks



1. CISS as measure of systemic stress

"You can't see the wood for the trees"



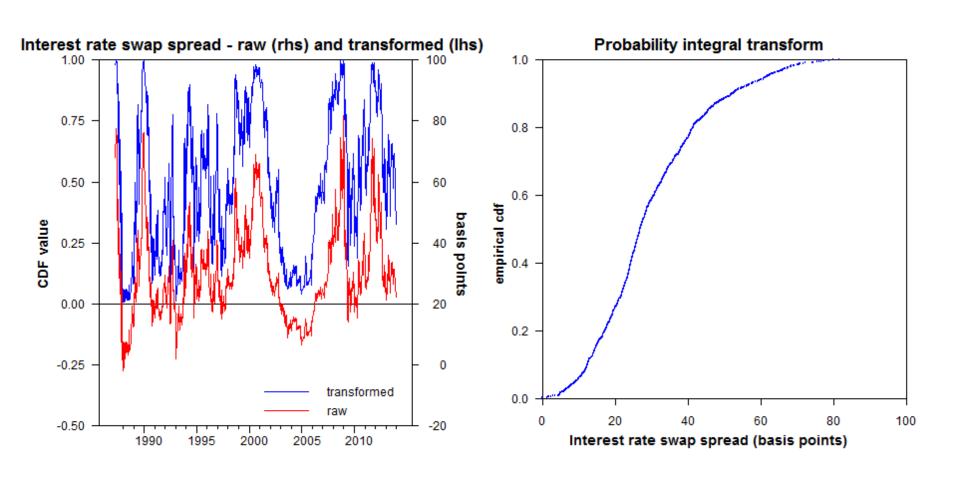
1. CISS as measure of systemic stress (cont'd)

- ECB's Composite Indicator of Systemic Stress (Hollo, Kremer and Lo Duca, 2012);
- Novel financial stress index which builds on standard definitions of systemic risk (focus on widespread financial instability);
- Identifies 5 key segments of financial system: 1) Financial intermediaries; 2) Money market; 3) Bond market; 4) Stock market;
 5) Foreign exchange market;
- Populates each of the 5 segments with 3 representative standard price-based stress indicators (incl. volatilities, spreads, cumulated valuation losses, etc.);
- Raw indicators transformed by empirical CDF (probability integral transform) to harmonise scale [0, 1] and distribution (std. uniform)
- Compute 5 subindexes of stress by taking arithmetic averages;
- Main innovation: aggregate subindexes based on the time-varying rank correlations between them (analogue to computation of portfolio risk from asset return variances and covariances)
 - Puts more weight on situations of widespread instability.



1. CISS – Data transformation

Probability integral transform: example case



1. CISS - Portfolio-theoretic aggregation framework

- Individual indicators aggregated into composite indicator based on portfolio theory
 - Compute system-wide stress analogous to portfolio risk in a static CAPM framework
 - Simple two-asset example:

$$\sigma_p^2 = w_1^2 \sigma_1^2 + w_2^2 \sigma_2^2 + 2w_1 w_2 \rho \sigma_1 \sigma_2$$

– General *n*-asset case:

$$\sigma_p^2 = \sum_{i=1}^n \sum_{j=1}^n w_i w_j \operatorname{cov}_{ij} = \sum_{i=1}^n \sum_{j=1}^n w_i w_j \rho_{ij} \sigma_i \sigma_j$$
$$= w' \sum w = (w \circ \sigma)' \Omega(w \circ \sigma)$$

 Σ a variance - covariance matrix and Ω a correlation matrix

1. CISS - Portfolio-theoretic aggregation framework

Applied to the present context:

- replace asset risk (σ_i) with transformed individual stress measures (s_i) ;
- portfolio risk (σ_p^2 or σ_p) then yields our composite stress index (CISS).

Two avenues to introduce systemic risk features:

- time-varying cross-correlations collected in correlation matrix Ω_t
 - ⇒ widespread instability, interconnectedness
- segment-specific "market shares" $w_{i,t}$ in the "portfolio" of stress measures $s_{i,t}$
 - ⇒ systemic importance (size, flows, real impact, ...).

1. CISS formula

$$CISS_{t} = (w_{t} \circ s_{t})'\Omega_{t}(w_{t} \circ s_{t}) \in (0,1] \quad \text{or}$$
$$= \sqrt{(w_{t} \circ s_{t})'\Omega_{t}(w_{t} \circ s_{t})}$$

$$w_{t}^{segments} = (w_{MM,t}, w_{BM,t}, w_{EM,t}, w_{FI,t}, w_{FX,t})'$$

$$= (\overline{w}_{MM}, \overline{w}_{BM}, \overline{w}_{EM}, \overline{w}_{FI}, \overline{w}_{FX})' = (0.19, 0.22, 0.14, 0.25, 0.20)'$$

$$s_{t} = (s_{MM,1,t}, s_{MM,2,t}, s_{MM,3,t}, s_{BM,1,t}, s_{BM,2,t}, s_{BM,3,t}, \cdots s_{FX,2,t}, s_{FX,3,t})' \text{ is a 1×15 vector}$$

$$w_{t} = 1/3 \cdot (\overline{w}_{MM}, \overline{w}_{MM}, \overline{w}_{MM}, \overline{w}_{BM}, \overline{w}_{BM}, \overline{w}_{BM}, \cdots \overline{w}_{FX}, \overline{w}_{FX}, \overline{w}_{FX})' \text{ is a 1×15 vector}$$

$$\Omega_{t} = \begin{pmatrix}
1 & \rho_{1,2,t} & \cdots & \rho_{1,15,t} \\
\rho_{12,t} & 1 & & \rho_{2,15,t} \\
\vdots & \cdots & \cdots & \vdots \\
\rho_{1,15,t} & \rho_{2,15,t} & \cdots & 1
\end{pmatrix} \text{ is a } 15 \times 15 \text{ matrix}$$



1. CISS - Time-varying cross-correlations

Time-varying cross-correlations

- computed as exponentially weighted moving averages (EWMA):
- time-varying version of Spearman's rank correlation.

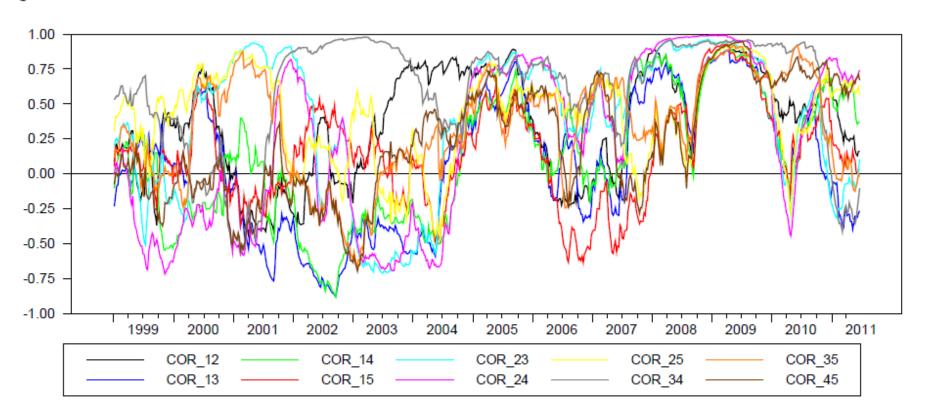
$$\sigma_{ij,t} = \lambda \sigma_{ij,t-1} + (1 - \lambda) \widetilde{s}_{i,t} \widetilde{s}_{j,t} \qquad \widetilde{s}_{i,t} = (s_{i,t} - 0.5)$$

$$\sigma_{i,t}^2 = \lambda \sigma_{i,t-1}^2 + (1 - \lambda) \widetilde{s}_{i,t}^2 \qquad \lambda = 0.93$$

$$\rho_{ij,t} = \sigma_{ij,t} / \sigma_{i,t} \sigma_{j,t}$$

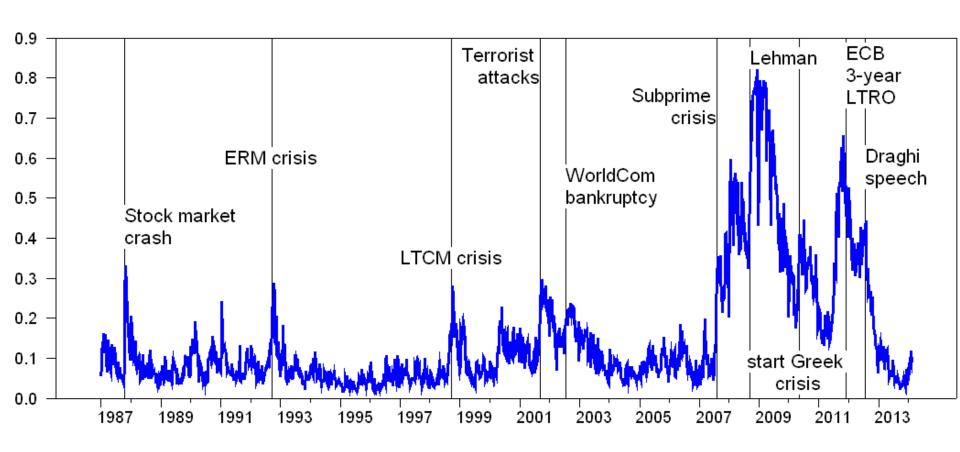
1. CISS - Time-varying cross-correlations

Fig. 2. Cross-correlations between subindices



Notes: Correlation pairs are computed as exponentially-weighted moving averages with smoothing parameter λ =0.93. The cross-correlations are labelled as follows: 1 – money market, 2 – bond market, 3 – equity market, 4 – financial intermediaries, 5 – foreign exchange market. Weekly euro area data from 8 Jan. 1999 to 24 June 2011.

1. CISS and financial stress events, 1987 to 2010



2. MS-BVAR model setup

- Endogenous variables: $y_t = [\Delta IP_t, \Delta P_t, R_t, \Delta Ln_t, S_t]$
 - ΔIP_t : growth rate of industrial production
 - ΔP_t : inflation rate (HICP)
 - R_t : money market rate (3-month Euribor)
 - $\Delta L n_t$: growth rate of bank lending
 - S_t : CISS
- Identification
 - Choleski decomposition, ordering of variables as above
 - Systemic financial stress responds instantaneously to innovations in all other variables (but not vice versa)
 - Results robust to other orderings
- Data sample: Euro area
 - monthly frequency, at annual rates
 - January 1987 to December 2010
 - seasonally adjusted

3. Determining macroeconomic regimes

Table 1:

Types of regime changes	no changes	variance change		coefficient change	variance and coefficient change	
Number of regimes	1v1c	2v1c	3v1c	1v2c	2v2c	3v2c
Log MDD	-6.05	92.36	131.95	37.76	126.08	147.36
Diff. to constant parameter model (1v1c)	0.00	98.41	138.00	43.81	132.13	153.41

Notes: Log Marginal data densities (mdd) are calculated as in Sims, Waggoner and Zha (2008); 1v1c: constant parameter model; ivjc: i shock variance regimes, j coefficient regimes.

Conclusions

- There are significant regime shifts
- Preferred specification has 3 variance and 2 coefficient regimes

3. Interpreting regimes: variance regimes

Table 2:

Regime label	production	inflation	interest rate	loans	CISS
1: v1	1.000	1.000	1.000	1.000	1.000
2: v2	0.905	1.525	0.285	0.738	0.617
3: v3	0.853	1.994	0.645	0.563	2.980

Notes: Standard deviation of the first variance regime (v1) normalised to one by construction.

Conclusions

- No uniform ranking order of shock volatilities across variables
- But regime v3 associated with very large financial stability shocks

3. Interpreting regimes: descriptive statistics

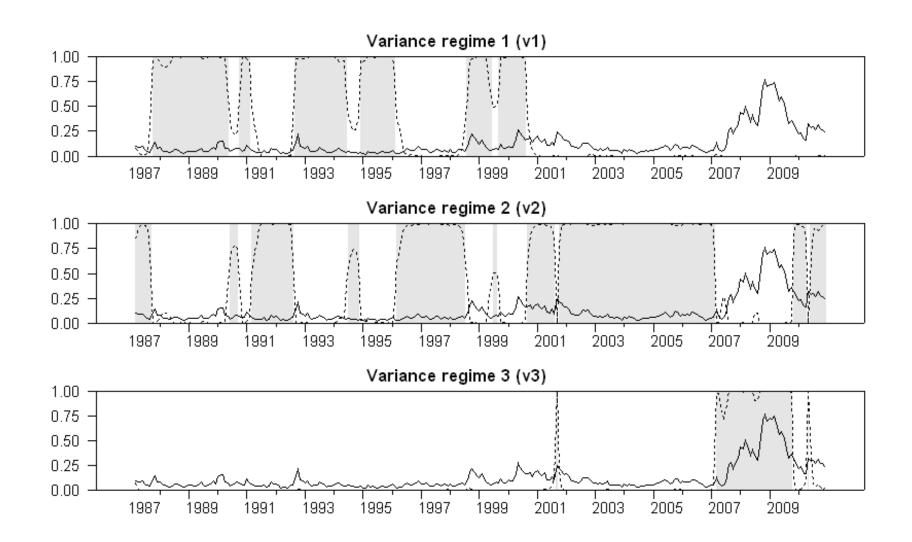
Table 3:

Regime label		sample shares				
	production	inflation	interest rate	loans	CISS	
1: v1, c1	0.535	2.264	5.850	5.973	0.071	16.1%
2: v1, c2	3.391	3.007	6.130	8.426	0.092	17.8%
3: v2, c1	2.783	1.959	3.219	6.325	0.081	35.3%
4: v2, c2	1.163	2.834	5.850	6.108	0.110	18.9%
5: v3, c1	3.958	2.430	4.178	9.658	0.260	5.2%
6: v3, c2	-11.293	1.571	2.878	4.655	0.520	6.6%

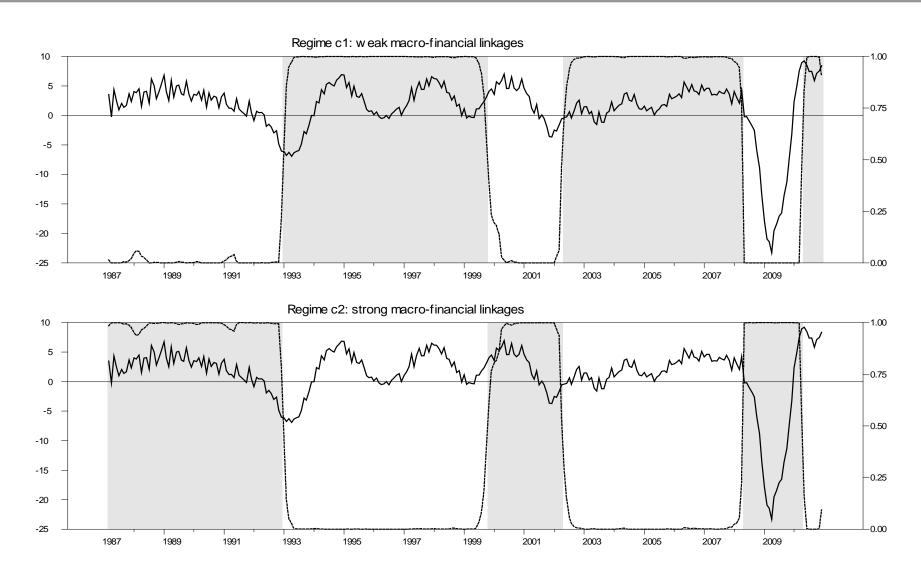
Notes: v=variance regime (1, 2, or 3); c=coefficient regime (1 or 2)

Regime 6 stands out: highest reading for CISS, lowest for all other variables

3. CISS and variance regime probabilities (dominant regimes shaded)

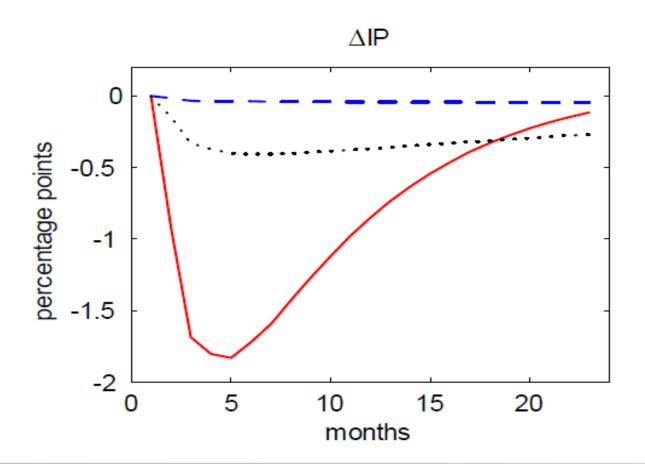


3. Industrial production growth and coefficient regime probabilities (dominant regimes shaded)



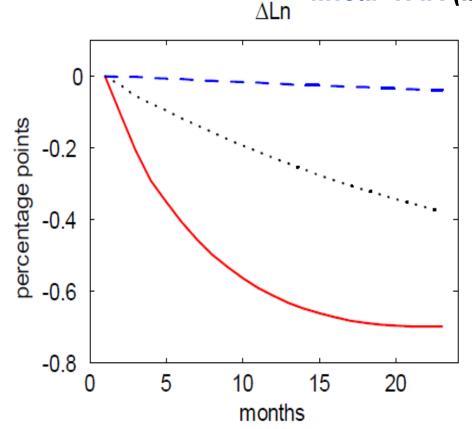
3. Impulse response functions: output growth to CISS shocks

"systemic crisis" regime 6 (red solid line)
"tranquil times" regime 1 (blue dashed line)
linear VAR (black dotted line)

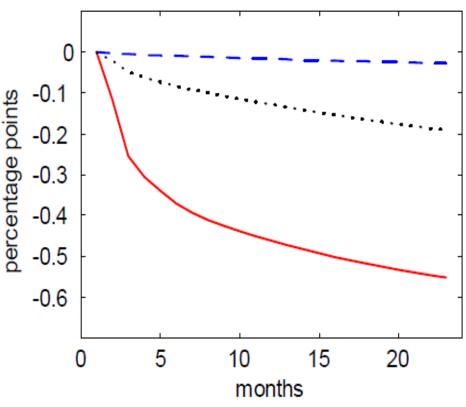


3. Impulse response functions: short rate and loans to CISS shocks

"systemic crisis" regime 6 (red line)
"tranquil times" regime 1 (blue dashed line)
linear VAR (black dotted line)

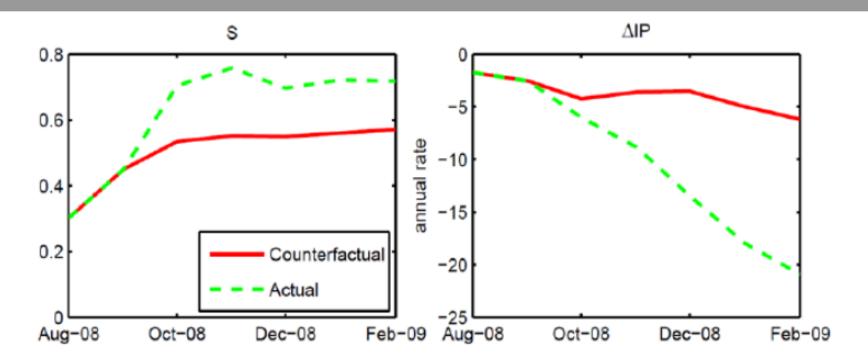






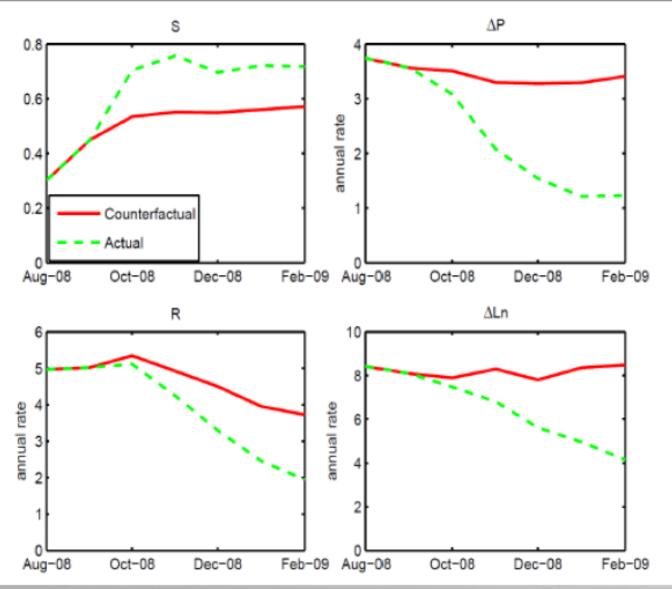
 Despite strong reaction of standard monetary policy

3. Counterfactual analysis I



- Counterfactual: Oct 2008 to Feb 2009, tranquil times (regime 1) instead of systemic crisis (regime 6)
 - Financial stress (S) at considerably lower levels
 - Decline in industrial production growth (△IP) much more muted

3. Counterfactual analysis I (cont'd)

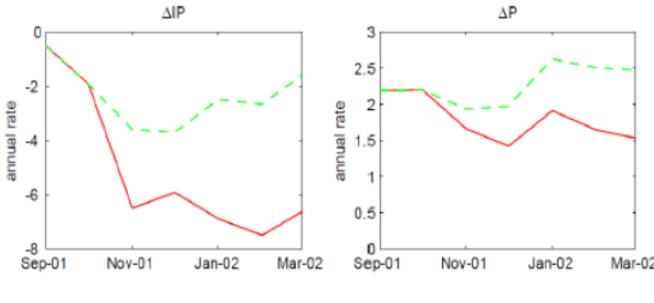


Counterfactual:

Oct 2008 to Feb 2009, tranquil times (regime 1) instead of systemic crisis (regime 6)

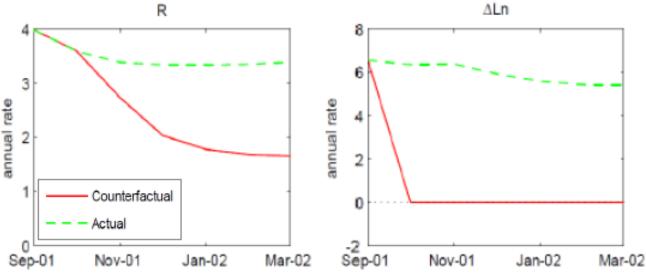
Also inflation
 (ΔP), monetary
 policy interest
 rate (R) and
 loan growth
 (ΔLn) much
 lower

3. Counterfactual analysis II



Counterfactual:

Oct 2001 to Mar 2002 ("dot-com bust"), stable path of CISS at average level in that period, Mar-02 path in loan growth to 0% as in crisis

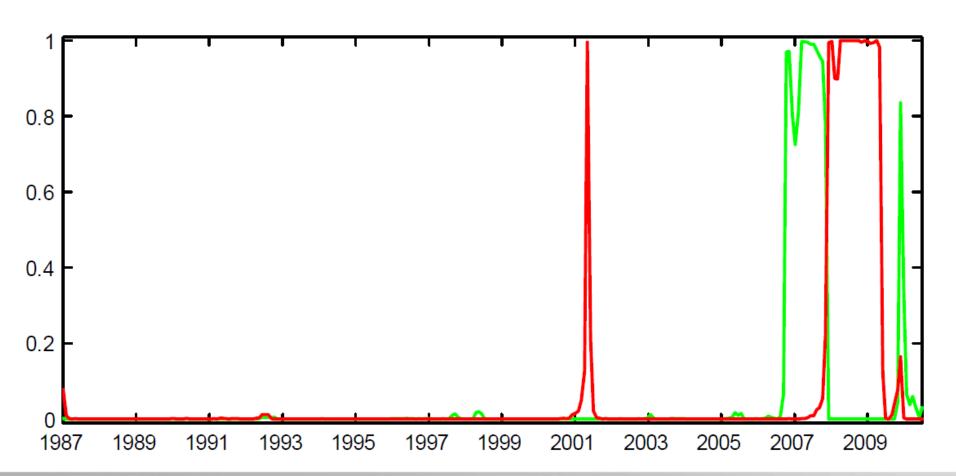


Output growth
 (∆IP), inflation
 (∆P) and short term interest
 rate (R) much
 lower

4. From subprime turmoil to systemic meltdown: A model tale

Smoothed regime probabilities

"systemic crisis" regime: high-stress variance + strong financial-real linkages "systemic fragility" regime: high-stress variance + weak financial-real linkages



5. Concluding remarks

- During systemic crisis the macroeconomic dynamics change fundamentally (evidence of non-linearities)
 - Both the parameters characterising the structure of the economy and the sizes of shocks switch regime;
 - Hence, the severe output losses typically caused by systemic crises may not only result from large shocks but also from the amplified transmission of shocks to the real economy.

Policy relevance

- Model may help explain large macroeconomic forecast errors during the recent crisis;
- Such MS-VAR models may improve macroeconomic nowcasting or even forecasting during episodes of severe systemic stress and may therefore represent a useful addition to the analytical toolkit of monetary and macro-prudential policy makers.

Annex

Markov-Switching Vector Autoregression model

$$y'_{t}A_{0}(s_{t}^{c}) = \sum_{l=1}^{p} y'_{t-l}A_{l}(s_{t}^{c}) + z'_{t}C(s_{t}^{c}) + \varepsilon'_{t}\Xi^{-1}(s_{t}^{v})$$
 (1)

 y_t : vector of endogenous variables [nx1];

 z_t : vector of exogenous variables and intercept terms; later assumed to be only intercepts [nx I];

 ε_t : error terms, vector of random shocks [nx1];

 Ξ : diagonal matrix containing the standard deviations of the shocks [nxn];

 A_0 [nxn], A_l [nxn], C [1xn]: coefficient matrices;

 s_t^c , s_t^v : unobserved state variables evolve according to two independent first-order Markov processes:

$$\Pr(s_t^m = i | s_{t-1}^m = k) = p_{ik}^m, \quad i, k = 1, 2, \dots h^m , \text{m=c,v}$$
 (2)

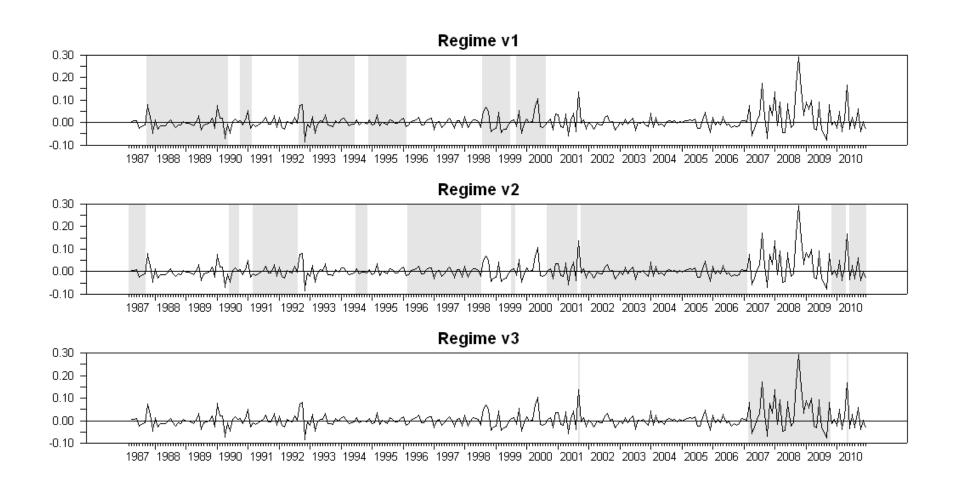
Let $Y^t = \{y_0, y_1, ... y_t\}$ as the vector y stacked in the time dimension, then the structural disturbances are conditionally normal:

$$\varepsilon'_t(s_t^v)|Y^{t-1}\sim N(0_{n\,x1},I_n) \tag{3}$$

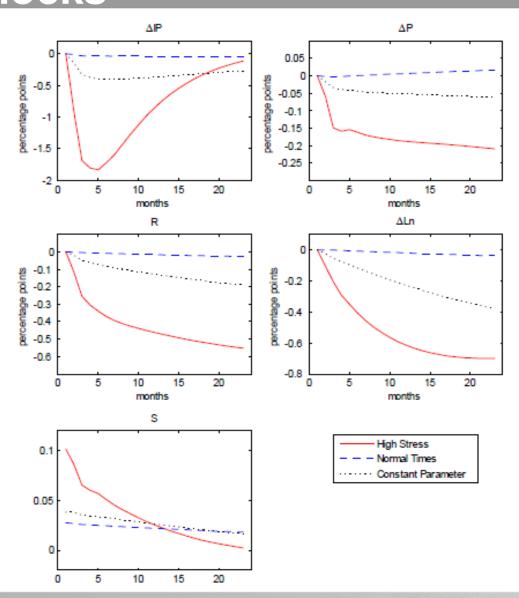
Model estimation and evaluation

- Bayesian methods generally following Sims, Waggoner and Zha (2008)
- Estimation of posterior mode ("most likely" estimate)
 - Blockwise optimisation algorithm
 - Parameters divided into blocks and initial guesses for the parameters used in a hill-climbing quasi-Newton optimisation procedure
 - Two sets of priors:
 - VAR coefficients: standard Minnesota priors
 - Transition matrix: Dirichlet prior
- Model evaluation: Marginal Data Densities (MDDs)
 - MDDs or posterior model probs are a measure of model fit (the higher, the better)
 - Computation according to Sims, Waggoner and Zha (2008)

CISS residuals and variance regimes (regimes shaded)

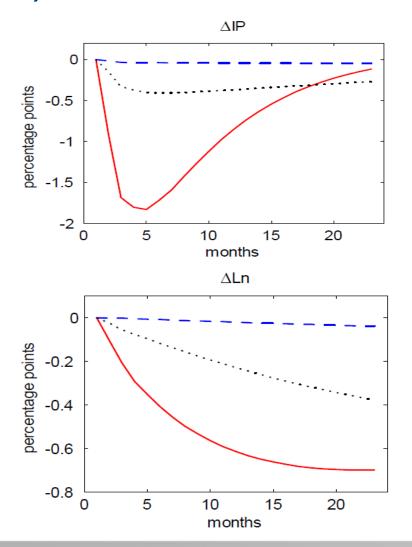


Full set of impulse response functions to CISS shocks



Alternative measures of financial stress: realised stock market volatility

1) IRFs with CISS



2) IRFs with stock market vola.

